Practice Final 1

- 1. Urn A has 4 blue balls. Urn B has 1 blue ball and 3 red balls.
 - (a) You draw a ball from a random urn and it is blue. What is the probability that it came from urn A?

Solution: Let B_1 be the event the ball is blue and A be the event the ball came from urn A. By Bayes' rule

$$P(A|B_1) = \frac{P(B_1|A) P(A)}{P(B_1|A) P(A) + P(B_1|A^c) P(A^c)} = \frac{1 \cdot (1/2)}{1 \cdot (1/2) + (1/4) \cdot (1/2)} = \frac{4}{5}.$$

(b) You draw another ball from the same urn. What is the probability that the second ball is also blue?

Solution: Let B_2 be the event that the second ball is blue. By the total probability theorem and Bayes' rule

$$P(B_2|B_1) = \frac{P(B_2 \cap B_1)}{P(B_1)} = \frac{P(B_2 \cap B_1|A) P(A) + P(B_2 \cap B_1|A^c) P(A^c)}{P(B_1|A) P(A) + P(B_1|A^c) P(A^c)} = \frac{1 \cdot (1/2) + (1/4)^2 \cdot (1/2)}{1 \cdot (1/2) + (1/4) \cdot (1/2)} = \frac{17}{20}$$

- 2. Computers A and B are linked through routers R_1 to R_4 as in the picture. Each router fails independently with probability 10%.
 - (a) What is the probability there is a connection between A and B?

Solution: Let R_i be the event that router *i* is operational. The event "there is a connection between *A* and *B*" is $(R_1 \cup R_2) \cap (R_3 \cup R_4)$. By independence

$$P((R_1 \cup R_2) \cap (R_3 \cup R_4)) = P(R_1 \cup R_2) P(R_3 \cup R_4)$$

= $(1 - P(R_1^c \cap R_2^c))(1 - P(R_1^c \cap R_2^c))$
= $(1 - P(R_1^c) P(R_2^c))(1 - P(R_3^c) P(R_4^c))$
= $(1 - 0.1^2)^2$
= $0.9801.$



(b) Are the events "there is a connection between A and B" and "exactly two routers fail" independent? Justify your answer.

- **Solution:** No. The probability that there is a connection between A And B given that exactly two routers fail is 2/3: Given that exactly two routers fail, the failed routers are equally likely to be any of the 6 pairs R_1R_3 , R_1R_4 , R_2R_3 , R_2R_4 , R_1R_2 , R_3R_4 , and there is a connection between A and B in the first 4 out of these 6 possibilities. This probability is not equal to the unconditional probability from part (a) and so the two events are not independent.
- 3. A bus takes you from A to B in 10 minutes. On average a bus comes once every 5 minutes. A taxi takes you in 5 minutes, and on average a taxi comes once every 10 minutes. Their arrival times are independent exponential random variables. A bus comes first.
 - (a) If you want to minimize the (expected) travel time, should you take this bus?

Solution: Yes. If you waited for a taxi your expected travel time would be the expected waiting time for the next taxi which is 10 minutes plus its travel time which is another 5 minutes for a total of 15 minutes.

(b) If you do take the bus, what is the probability that you made the wrong decision?

Solution: The probability of a wrong decision is the probability that a taxi arrives within the next five minutes, which is the probability that an Exponential(1/10) random variable is less than 5, which is $1 - e^{-5/10} = 1 - e^{-1/2} \approx 39.35\%$.

- 4. 10 people toss their hats and each person randomly picks one. The experiment is repeated one more time.
 - (a) What is the probability that Bob picked his own hat both times?

Solution: By independence, the probability that Bob picked his hat both times is the product of the probabilities that he picked it in each trial, so it is $(1/10) \cdot (1/10) = 1/100$.

(b) Let A be the event that at least one person picked their own hat both times. True or false: P(A) > 25%? Justify your answer.

Solution: False. Let X_i take value 1 if person *i* picked their hat both times. A occurs if $X = X_1 + \cdots + X_{10} \ge 0$. By part (a) and linearity of expectation, $E[X] = 10 \cdot (1/100) = 0.1$. By Markov's inequality, $P[X \ge 1] \le E[X]/1 = 0.1$ which is less than 25%.

- 5. X is a Normal $(0, \Theta)$ random variable, where the prior PMF of the parameter Θ is $P(\Theta = 1/2) = 1/2$, $P(\Theta = 1) = 1/2$. You observe the following three independent samples of X: 1.0, 1.0, -1.0.
 - (a) What is the posterior PMF of Θ ?

Solution: By Bayes' rule

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$$f_{\Theta|X_1X_2X_3}(\theta|1.0, 1.0, -1.0) \propto f_{X_1X_2X_3|\Theta}(1.0, 1.0, -1.0|\theta) \operatorname{P}(\Theta = \theta) \propto \frac{1}{\theta^3} e^{-3/2\theta^2} \operatorname{P}(\Theta = \theta).$$

As Θ is equally likely to take values 0 and 1/2, the posterior PMF is

$$f_{\Theta|X_1X_2X_3}(1/2|1.0, 1.0, -1.0) = \frac{8e^{-6}}{8e^{-6} + e^{-3/2}} \quad f_{\Theta|X_1X_2X_3}(1|1.0, 1.0, -1.0) = \frac{e^{-3/2}}{8e^{-6} + e^{-3/2}}.$$

(b) What is the MAP estimate of Θ ?

Solution: As $e^{-3/2} \approx 0.2231$ is larger than $8e^{-6} \approx 0.0198$ the MAP estimate is $\hat{\Theta} = 1$.

(c) What is the posterior probability that $|X| \ge 1$?

Solution: The posterior probabilities of Θ are 1/2 with probability about $0.0198/(0.2231 + 0.0198) \approx 0.0815$ and 1 with probability about $0.2231/(0.2231 + 0.0198) \approx 0.9185$. By the total probability theorem the posterior probability of $|X| \ge 1$ is about

$$\begin{aligned} 0.0185 \cdot P(|Normal(0, 1/2)| \geq 1) + 0.9185 P(|Normal(0, 1)| \geq 1) \\ \approx 0.0185 \cdot 2 P(Normal(0, 1) \geq 2) + 0.9185 \cdot 2 P(Normal(0, 1) \geq 1) \\ \approx 0.0185 \cdot 2 \cdot 0.023 + 0.9185 \cdot 2 \cdot 0.159 \\ \approx 0.2929. \end{aligned}$$

Practice Final 2

- 1. Let X, Y, Z be independent Binomial $(2, \frac{1}{2})$ random variables.
 - (a) What is the conditional PMF of X conditioned on $X \neq Z$?

Solution: The joint PMF is

$$P(X = 0, X \neq Z) = P(X = 0, Z = 1) + P(X = 0, Z = 2) = \frac{1}{4} \cdot \frac{1}{2} + \frac{1}{4} \cdot \frac{1}{4} = \frac{3}{16}$$

$$P(X = 1, X \neq Z) = P(X = 1, Z = 0) + P(X = 1, Z = 2) = \frac{1}{2} \cdot \frac{1}{4} + \frac{1}{2} \cdot \frac{1}{4} = \frac{4}{16}$$

$$P(X = 2, X \neq Z) = P(X = 2, Z = 0) + P(X = 2, Z = 1) = \frac{1}{4} \cdot \frac{1}{4} + \frac{1}{4} \cdot \frac{1}{2} = \frac{3}{16}$$

The conditional PMF is the joint PMF normalized by $P(X \neq Z)$, which is

$$\mathbf{P}(X=0|X\neq Z) = \frac{3}{10}, \quad \mathbf{P}(X=1|X\neq Z) = \frac{4}{10}, \quad \mathbf{P}(X=2|X\neq Z) = \frac{3}{10}$$

(b) Are X and Y independent conditioned on $(X \neq Z)$ AND $(Y \neq Z)$?

Solution: No. We show that

$$P(X = 1 \mid X \neq Z, Y \neq Z) P(Y = 1 \mid X \neq Z, Y \neq Z) \neq P(X = 1, Y = 1 \mid X \neq Z, Y \neq Z).$$

By symmetry the two probabilities on the left are the same. We calculate these expressions:

$$\begin{split} \mathrm{P}(X \neq Z, Y \neq Z) &= \sum_{z \in \{0,1,2\}} \mathrm{P}(X \neq Z, Y \neq Z | Z = z) \, \mathrm{P}(Z = z) \\ &= (3/4)^2 \cdot (1/4) + (1/2)^2 \cdot (1/2) + (3/4)^2 \cdot (1/4) \\ &= 13/32, \\ \mathrm{P}(X = 1, X \neq Z, Y \neq Z) &= \mathrm{P}(X = 1, Z \neq 1, Y \neq Z) \\ &= \mathrm{P}(X = 1) \, \mathrm{P}(Z \neq 1, Y \neq Z) \\ &= (1/2) \cdot (\mathrm{P}(Y \neq 0, Z = 0) + \mathrm{P}(Y \neq 2, Z = 2)) \\ &= (1/2) \cdot 2 \cdot (1/4) \cdot (3/4) \\ &= 3/16, \\ \mathrm{P}(X = 1, Y = 1, X \neq Z, Y \neq Z) &= \mathrm{P}(X = 1, Y = 1, Z \neq 1) \\ &= (1/2)(1/2)(1/2) \\ &= 1/8. \end{split}$$

By the conditional probability formula the expression on the left is $((3/16)/(13/32))^2 \approx 0.2130$ and the one on the right is $(1/8)/(13/32) \approx 0.3077$. These are not equal.

- 2. Alice and Bob decide to meet somewhere. Alice's arrival time A is uniform between 12:00 and 12:45. Bob's arrival time B is uniform between 12:15 and 1:00. Their arrival times are independent.
 - (a) Let f_{A-B} be the PDF of A B. What is $f_{A-B}(0)$?

Solution: We model A and B as Uniform(0, 3/4) and Uniform(1/4, 1) random variables respectively (at the hour scale). By the convolution formula, $f_{A-B}(0) = \int_{-\infty}^{\infty} f_A(t) f_B(t) dt$, where f_A , f_B are the PDFs of A and B. $f_A(t) f_B(t)$ takes value $(4/3)^2$ when t is between 1/4 and 3/4 and 0 otherwise, so the integral equals $(1/2) \cdot (4/3)^2 = 8/9$. (If time is scaled in minutes the answer is 60 times smaller.)

(b) What is the probability that Bob arrives before Alice?

Solution: The event that Bob arrives before Alice is the value of the integral $\int_{a>b} f_A(a) f_B(b) dadb$. The value of the integrand is $(4/3)^2$ when (a, b) is in the interior of the triangle with vertices (1/4, 1/4), (1/4, 3/4), (3/4, 3/4) and zero elsewhere. The triangle has area $(1/2)^2/2 = 1/8$. Therefore $P(A > B) = (1/8)(4/3)^2 = 2/9$.

- 3. Let Y = AX + B where A, B, X are independent Normal(0, 1) random variables.
 - (a) What is $\operatorname{Var}[\operatorname{E}[Y|X]]$?

Solution: By linearity of expectation, E[AX+B|X] = E[A]X+E[B] = 0 so Var[E[Y|X]] = 0.

(b) What is E[Var[Y|X]]?

Solution: By independence, $\operatorname{Var}[AX + B|X] = \operatorname{Var}[AX|X] + \operatorname{Var}[B] = X^2 \operatorname{Var}[A] + \operatorname{Var}[B] = X^2 + 1$, so $\operatorname{E}[\operatorname{Var}[Y|X]] = \operatorname{E}[X^2 + 1] = \operatorname{Var}[X] + 1 = 2$.

4. Boys and girls arrive independently at a meeting point at a rate of one boy per minute and one girl per minute, respectively. Let T be the first time at which both a boy and a girl have arrived.

(a) Find the cumulative distribution function (CDF) of T.

Solution: The probability that a boy has arrived by time t is $1 - e^{-t}$, i.e. the CDF of an Exponential(1) random variable. The probability that a boy has arrived by time t is therefore $1 - e^{-t}$, and same for a girl. The events are independent, the probability that both have arrived by time t is $P(T \le t) = (1 - e^{-t})^2$ if $t \ge 0$ and 0 if not.

(b) What is the expected value of T? (**Hint:** You don't *have* to use calculus.)

Solution: We can write $T = T_1 + T_2$ where T_1 is the arrival time of the first person and T_2 is the arrival time of the next person of the opposite gender. As people arrive at a rate of two per minute, T_1 is an Exponential(2) random variable. By the memoryless property T_2 is an Exponential(1) random variable. Therefore $E[T] = E[T_1] + E[T_2] = 1/2 + 1 = 3/2$.

- 5. A deck of cards is divided into 26 pairs. Let X be the number of those pairs in which both cards are of the same suit. (A deck of cards has 4 suits and each suit has 13 cards.)
 - (a) What is the expected value of X?

Solution: We can write $X = X_1 + \cdots + X_{26}$ where X_i is 1 if the cards in the *i*-th pair are of the same suit and 0 if not. Then $E[X_i] = P(X_i = 1)$ is the probability that the *i*-th pair's cards are of the same suit, which is 12/51 because conditioned on the first card's suit, there are 12 out of 51 identical choices for the second one. By linearity of expectation $E[X] = E[X_1] + \cdots + E[X_{26}] = 26 \cdot 12/51 \approx 6.118.$

(b) What is the variance of X?

Solution: The variance of X is the sum of the 26 variances of X_i and the $26 \cdot 25$ covariances of X_i and X_j . The variance of X_i is $v = (12/51) \cdot (1 - 12/51) \approx 0.1799$. The covariance of X_i and X_j is

$$E[X_i X_j] - E[X_i] E[X_j] = P(X_i = 1, X_j = 1) - P(X_i = 1) P(X_j = 1)$$

The term $P(X_i = 1, X_j = 1)$ is the probability of the event A that within both the *i*-th pair and the *j*-th pair, both cards are of the same suit. We can calculate this using the total probability theorem applied to the event E that the first card of the *i*-th pair and the first card of the *j*-th pair are of the same suit:

$$P(X_i = 1, X_j = 1) = P(A) = P(A|E) P(E) + P(A|E^c) P(E^c)$$

The probability of E is 12/51. Conditioned on E, A happens if the second cards of both pairs are also of the same suit, which is $11/50 \cdot 10/49$. Conditioned on E^c —for example, if the *i*-th pair's first card is a heart and the *j*-th pair first card is a spade—A happens if the second cards are a heart and a spade respectively, which happens with probability $(12/50) \cdot (12/49)$, and so

$$P(A) = \frac{11}{50} \cdot \frac{10}{49} \cdot \frac{12}{51} + \frac{12}{50} \cdot \frac{12}{49} \cdot \left(1 - \frac{12}{51}\right).$$

Therefore the covariance of X_i and X_j equals

$$c = P(A) - \left(\frac{12}{51}\right)^2 \approx 0.0001469.$$

Finally, $Var[X] = 26 \cdot v + 26 \cdot 25 \cdot c \approx 4.7737.$