

Practice questions

1. You flip a coin with unknown probability of heads Θ . Assume a Uniform(0, 1) prior on Θ .

(a) You observe the outcome HTH after three flips. What is the posterior probability that the next flip will be a head?

Solution: The posterior Θ given two heads and a tail is a Beta(3, 2) random variable with PDF $f_{\Theta|X_1X_2X_3}(\theta|\text{HTH}) = 12\theta^2(1 - \theta)$. The probability that the next toss is a head given $\Theta = \theta$ is $P(X_4 = \text{H}|\Theta = \theta) = \theta$. By the total probability theorem

$$P(X_4 = \text{H}) = \int_0^1 P(X_4 = \text{H}|\Theta = \theta)f_{\Theta|X_1X_2X_3}(\theta|\text{HTH})d\theta = \int_0^1 \theta \cdot 12\theta^2(1 - \theta)d\theta = \frac{3}{5}.$$

(b) (**Optional**) More generally, you observe h heads and $t = n - h$ tails after n flips. What is the posterior probability that the next flip will be a head?

Solution: By the same reasoning, the posterior on Θ is Beta($h + 1, t + 1$), and the probability that the next toss is a head given $\Theta = \theta$ is θ . The posterior probability that the next flip is a head is

$$\begin{aligned} P(X_4 = \text{H}) &= \int_0^1 P(X_{n+1} = \text{H}|\Theta = \theta)f_{\Theta|X_1 \dots X_n}(\theta|h \text{ heads}, t \text{ tails})d\theta \\ &= \int_0^1 \frac{1}{B(h+1, t+1)} \cdot \theta^h(1 - \theta)^t \cdot \theta d\theta \\ &= \frac{B(h+2, t+1)}{B(h+1, t+1)} \cdot \int_0^1 \frac{1}{B(h+2, t+1)} \cdot \theta^{h+1}(1 - \theta)^t \cdot \theta d\theta \\ &= \frac{B(h+2, t+1)}{B(h+1, t+1)}. \end{aligned}$$

The last equality is true because the integrand is the PDF of a Beta($h + 2, t + 1$) random variable so the integral evaluates to one. Using the formula for the Beta function we get that

$$P(X_4 = \text{H}) = \frac{(h+1)!t!/(h+t+2)!}{h!t!/(h+t+1)!} = \frac{h+1}{h+t+2} = \frac{h+1}{n+2}.$$

2. Romeo has figured out that girls have a hidden parameter Θ that determines how late they tend to show up to dates. As usual his prior on Θ is Uniform(0, 1).

(a) Assume that that likelihood $f_{X|\Theta}$ of Juliet being X hours late on a given date is Uniform(0, Θ). Calculate the probability that Juliet shows up at least 10 minutes late at her first date.

Solution: The likelihood of X is $f_{X|\Theta}(x|\theta) = 1/\theta$ when $0 \leq x \leq \theta$ and zero otherwise. The conditional probability of the event $X \geq 1/6$ given $\Theta = \theta$ is

$$P(X \geq 1/6|\Theta = \theta) = \int_{1/6}^{\infty} f_{X|\Theta}(x|\theta)dx = \int_{1/6}^{\theta} \frac{1}{\theta} dx = \begin{cases} (1/\theta)(\theta - 1/6), & \text{if } \theta \geq 1/6 \\ 0, & \text{otherwise.} \end{cases}$$

The total probability is

$$P(X \geq 1/6) = \int_{-\infty}^{\infty} P(X \geq 1/6|\Theta = \theta)f_{\Theta}(\theta)d\theta = \int_{1/6}^1 \frac{1}{\theta}(\theta - 1/6) \cdot 1d\theta = \frac{5 - \ln 6}{6},$$

which is about 0.535.

- (b) Juliet shows up exactly 10 minutes late at her first date. What is the probability that she shows up at least 10 minutes late at her second date?

Solution: Let X' be Juliet's arrival time at her second date. The likelihood of X' is the same as the likelihood of X , but the PDF of Θ is now the posterior PDF after the first date, namely

$$f_{\Theta|X}(\theta|1/6) \propto f_{X|\Theta}(1/6|\theta)f_{\Theta}(\theta) = \begin{cases} 1/\theta, & \text{if } 1/6 \leq \theta \leq 1 \\ 0, & \text{otherwise.} \end{cases}$$

The proportionality constant must be $1/\int_{1/6}^1(1/\theta)d\theta = 1/\ln 6$. The total probability is therefore

$$\begin{aligned} P(X' \geq 1/6|X = 1/6) &= \int_{-\infty}^{\infty} P(X' \geq 1/6|\Theta = \theta)f_{\Theta|X}(\theta|1/6) \\ &= \int_{1/6}^1 \frac{1}{\theta}(\theta - 1/6) \cdot \frac{1}{(\ln 6)\theta} d\theta \\ &= 1 - \frac{5}{6 \ln 6}, \end{aligned}$$

which is about 0.535. The probability didn't change up to third decimal (but it is not the same).

- (c) Repeat parts (a) and (b) with an Exponential(Θ) likelihood.

Solution: For part (a) we would get that $P(X \geq 1/6|\Theta = \theta) = e^{-\theta/6}$ for $\theta \geq 0$. As the prior remains the same, $P(X \geq 1/6) = \int_0^1 e^{-\theta/6} d\theta = 6(1 - e^{-1/6}) \approx 0.921$. In Homework 2 the posterior PDF for Θ given $X = 1/6$ was calculated to be $f_{\Theta|X}(\theta|1/6) \approx 2.233\theta e^{-\theta/6}$ for $0 \leq \theta \leq 1$. Therefore the answer in part (b) changes to

$$P(X' \geq 1/6|X = 1/6) \approx \int_0^1 e^{-\theta/6} \cdot 2.233\theta e^{-\theta/6} d\theta \approx 0.897.$$

The probability decreased by a little bit.

3. In a group of ten people, including Alice and Bob, each pair is friends with probability P , independently of the other pairs.¹

- (a) Let A be the number of Alice's friends and B be the number of Bob's friends in the group. Conditioned on $P = p$, what kind of random variables are A and B ? Are they independent?

Solution: The PMF of A conditioned on P is:

$$P_{A|P}(A = a|P = p) = \binom{10-1}{a} p^a (1-p)^{10-1-a}$$

Therefore A is a binomial random variable and so is B . However they are not independent conditioned on P : $P_{A,B|P}(A = 9, B = 0|P = p) = 0$ (Alice is a friend of all other nine people, contradicting the fact that Bob has no friend) while $P(A = 9|P = p) \cdot P(B = 0|P = p) > 0$.

¹This is the Erdős-Rényi random graph model.

- (b) Now suppose P is unknown. Alice counts five friends in the group. What is her MAP estimate of P assuming a Uniform(0, 1) prior?

Solution: P given $A = 5$ is a Beta(5 + 1, 4 + 1) random variable. The MAP estimate for such a random variable is $5/(5 + 4) = 5/9$.

- (c) Bob, who is one of Alice's friends, tells her that he has only one other friend in the group. How does this information affect Alice's MAP estimate of P ?

Solution: Even though A and B are dependent, the random variable C that counts the total number of friendships that involve Alice *or* Bob is a Binomial(17, P) random variable, so P conditioned on $C = 5 + 1 = 6$ is a Beta(6 + 1, 11 + 1) random variable. The MAP estimate is 6/17.

4. The ENGG2780A TAs want to estimate the hardness of the problem that they prepared for an upcoming homework. They know from experience that the time X (in minutes) it takes them to solve the problem is Exponential(1/15) if the problem is easy and Exponential(1/25) if the problem is hard (like this one). Based on previous quizzes, the prior probability that the problem is hard is 0.3.

- (a) Fanbin solves the problem in 20 minutes. According to the MAP rule, is the problem easy or hard? What is the probability of error?

Solution: Let Θ be the difficulty of the problem. The problem is difficult if $\Theta = 1$, otherwise $\Theta = 0$.

$$P(\Theta = 1|X = 20) \propto f_{X|\Theta}(X = 20|\Theta = 1) P(\Theta = 1) = \frac{1}{25} e^{-20/25} \cdot 0.3 \approx 0.00539$$

$$P(\Theta = 0|X = 20) \propto f_{X|\Theta}(X = 20|\Theta = 0) P(\Theta = 0) \propto \frac{1}{15} e^{-20/15} \cdot 0.7 \approx 0.01230$$

The MAP rule predicts that the problem is simple ($MAP = 0$). The probability of error is

$$P(MAP \neq \Theta|X = 20) = P(\Theta = 1|X = 20) \approx \frac{0.00539}{0.00539 + 0.01230} \approx 0.305.$$

- (b) The other five TAs try it out and their recorded solution times are 10, 10, 15, 25, and 35 minutes, respectively. How do the answers in part (a) change? [Adapted from Bertsekas-Tsitsiklis problem 8.2.6]

Solution: Let X_1, X_2, \dots, X_5 be the solution time of the five TAs.

$$P(\Theta = 1|X_1 = 20, X_2 = 10, \dots, X_6 = 35) \propto f_{X_1|\Theta}(20|1) \cdots P_{X_6|\Theta}(35|1) P(\Theta = 1)$$

$$\propto \frac{1}{25^6} e^{-(20+10+10+15+25+35)/25} \cdot 0.3$$

$$\approx 1.235 \cdot 10^{-11}$$

Similarly,

$$P(\Theta = 0|X_1 = 20, X_2 = 10, \dots, X_6 = 35) \propto f_{X_1|\Theta}(20|0) \cdots P_{X_6|\Theta}(35|0) P(\Theta = 0)$$

$$\propto \frac{1}{15^6} e^{-(20+10+10+15+25+35)/15} \cdot 0.7$$

$$\approx 2.870 \cdot 10^{-11}$$

Again the MAP rule predicts the problem is easy, and the probability of error is about $1.235/(1.235 + 2.870) \approx 0.300$.